

CS 2750 Machine Learning

Lecture 15

Learning Bayesian belief networks

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Learning of BBN

Learning.

- **Learning of parameters of conditional probabilities**
- **Learning of the network structure**

Variables:

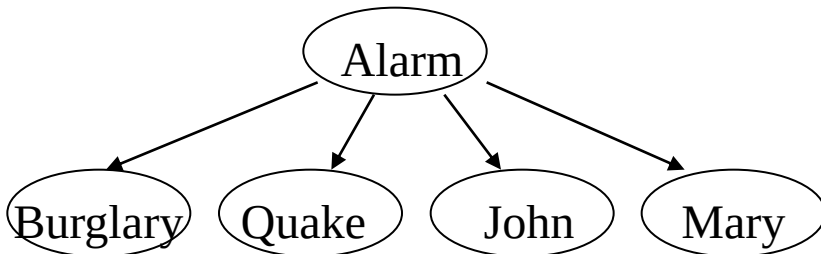
- **Observable** – values present in every data sample
- **Hidden** – they values are never observed in data
- **Missing values** – values sometimes present, sometimes not

Already Covered: Learning of parameters of BBN

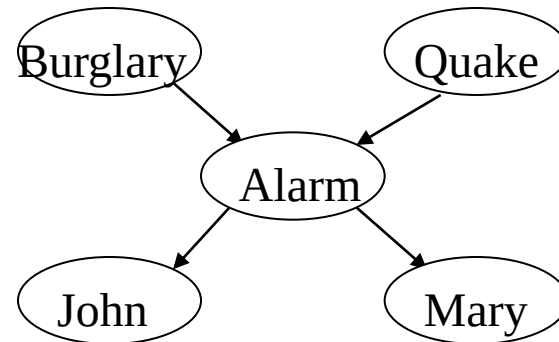
1. With observable variables
2. Hidden variables and missing values

Model selection

- **BBN has two components:**
 - **Structure of the network** (models conditional independences)
 - **A set of parameters** (conditional child-parent distributions)
- How to learn the structure?



?



Learning the structure

Criteria we can choose to score the structure S

- **Marginal likelihood**

$$\text{maximize } P(D | S, \xi)$$

ξ - represents the prior knowledge

- **Maximum posterior probability**

$$\text{maximize } P(S | D, \xi)$$

$$P(S | D, \xi) = \frac{P(D | S, \xi)P(S | \xi)}{P(D | \xi)}$$

How to compute marginal likelihood $P(D | S, \xi)$?

Learning of BBNs

- **Notation:**

- i ranges over all possible variables $i=1,\dots,n$
- $j=1,\dots,q$ ranges over all possible parent combinations
- $k=1,\dots,r$ ranges over all possible variable values

- Θ - parameters of the BBN

Θ_{ij} is a vector of Θ_{ijk} representing parameters of the conditional probability distribution; such that $\sum_{k=1}^r \Theta_{ijk} = 1$

N_{ijk} - a number of instances in the dataset where parents of variable X_i take on values j and X_i has value k

$$N_{ij} = \sum_{k=1}^r N_{ijk}$$

α_{ijk} - prior counts (parameters of Beta and Dirichlet priors)

$$\alpha_{ij} = \sum_{k=1}^r \alpha_{ijk}$$

Marginal likelihood

- Integrate over all possible parameter settings

$$P(D | S, \xi) = \int_{\Theta} P(D | S, \Theta, \xi) p(\Theta | S, \xi) d\Theta$$

- Using the assumption of parameter and sample independence

$$P(D | S, \xi) = \prod_{i=1}^n \prod_{j=1}^{q_i} \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} \prod_{k=1}^{r_i} \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})}$$

- We can use **log-likelihood score** instead

$$\log P(D | S, \xi) = \sum_{i=1}^n \left\{ \sum_{j=1}^{q_i} \log \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} + \sum_{k=1}^{r_i} \log \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})} \right\}$$

Score is decomposable along variables !!!

Appendix: marginal likelihood

- **Marginal likelihood**

$$P(D | S, \xi) = \int_{\Theta} P(D | S, \Theta, \xi) p(\Theta | S, \xi) d\Theta$$

- **From the iid assumption:**

$$P(D | S, \Theta) = \prod_{h=1}^N \prod_{i=1}^n P(x_i^h | \text{parents}_i^h, \Theta)$$

- Let r_i = number of values that attribute x_i can take
 q_i = number of possible parent combinations
 N_{ijk} = number of cases in D where x_i has value k and parents with values j .

$$\begin{aligned} &= \prod_i^n \prod_j^{q_i} \prod_k^{r_i} P(x_i = k | \text{parents}_i = j, \Theta)^{N_{ijk}} \\ &= \prod_i^n \prod_j^{q_i} \prod_k^{r_i} \theta_{ijk}^{N_{ijk}} \end{aligned}$$

Appendix: the marginal likelihood

$$P(D \mid S, \xi) = \int_{\Theta} P(D \mid S, \Theta, \xi) p(\Theta \mid S, \xi) d\Theta$$

- From parameter independence

$$p(\Theta \mid S, \xi) = \prod_{i=1}^n \prod_{j=1}^{q_i} p(\Theta_{ij} \mid S, \xi)$$

- Priors for $p(\Theta_{ij} \mid S, \xi)$

- $\Theta_{ij} = (\Theta_{ij1}, \dots, \Theta_{ijr_i})$ is a vector of parameters;
- we use a Dirichlet distribution with parameters α

$$P(\Theta_{ij} \mid S, \xi) = P(\Theta_{ij1}, \dots, \Theta_{ijr_i} \mid S, \xi) = \text{Dir}(\Theta_{ij1}, \dots, \Theta_{ijr_i} \mid \alpha)$$

$$= \frac{\Gamma(\sum_{k=1}^{r_i} \alpha_{ijk})}{\prod_{k=1}^{r_i} \Gamma(\alpha_{ijk})} \prod_{k=1}^{r_i} \Theta_{ijk}^{\alpha_{ijk} - 1}$$

Appendix: marginal likelihood

- Combine things together:

$$\begin{aligned} P(D | S_i) &= \int_{\Theta} P(D | S_i, \Theta) P(\Theta | S_i) d\Theta \\ &= \int \prod_i^n \prod_j^{q_i} \prod_k^{r_i} \Theta^{N_{ijk}} \cdot \frac{\Gamma(\sum_{k=1}^{r_i} \alpha_{ijk})}{\prod_{k=1}^{r_i} \Gamma(\alpha_{ijk})} \prod_{k=1}^{r_i} \Theta^{\alpha_{ijk} - 1} d\Theta \\ &= \prod_i^n \prod_j^{q_i} \frac{\Gamma(\sum_{k=1}^{r_i} \alpha_{ijk})}{\prod_{k=1}^{r_i} \Gamma(\alpha_{ijk})} \int \prod_{k=1}^{r_i} \Theta^{N_{ijk} + \alpha_{ijk} - 1} d\Theta \\ &= \prod_i^n \prod_j^{q_i} \frac{\Gamma(\alpha_{ij})}{\prod_{k=1}^{r_i} \Gamma(\alpha_{ijk})} \cdot \frac{\prod_{k=1}^{r_i} \Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ij} + N_{ij})} \end{aligned}$$

A trick to compute the marginal likelihood

- Integrate over all possible parameter settings

$$P(D | S, \xi) = \int_{\Theta} P(D | S, \Theta, \xi) p(\Theta | S, \xi) d\Theta$$

- Posterior of parameters, given data and the structure

$$p(\Theta | D, S, \xi) = \frac{P(D | \Theta, S, \xi) p(\Theta | S, \xi)}{P(D | S, \xi)}$$

Trick

$$P(D | S, \xi) = \frac{P(D | \Theta, S, \xi) p(\Theta | S, \xi)}{p(\Theta | D, S, \xi)}$$

- Gives the solution

$$P(D | S, \xi) = \prod_{i=1}^n \prod_{j=1}^{q_i} \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} \prod_{k=1}^{r_i} \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})}$$

Learning the structure

- **Likelihood of data for the BBN** (structure and parameters)

$$P(D | S, \Theta, \xi)$$

measures the goodness of fit of the BBN to data

- **Marginal likelihood** (for the structure only)

$$P(D | S, \xi)$$

- **Does not measure only a goodness of fit. It is:**
 - different for structures of different complexity
 - Incorporates preferences towards simpler structures, **implements Occam's razor !!!!**

Occam's Razor

- Why there is a preference towards simpler structures ?

Rewrite marginal likelihood as

$$P(D | S, \xi) = \frac{\int_{\Theta} P(D | S, \Theta, \xi) p(\Theta | S, \xi) d\Theta}{\int_{\Theta} p(\Theta | S, \xi) d\Theta}$$

We know that $\int_{\Theta} p(\Theta | S, \xi) d\Theta = 1$

Interpretation: in more complex structures there are more ways parameters can be set badly

- **The numerator:** count of good assignments
- **The denominator:** count of all assignments

Approximations of probabilistic scores

Approximations of the marginal likelihood and posterior scores

- **Information based measures**

- Akaike criterion
 - Bayesian information criterion (BIC)
 - Minimum description length (MDL)
- Reflect the tradeoff between the fit to data and preference towards simpler structures

Example: **Akaike criterion.**

Maximize: $score(S) = \log P(D | S, \Theta_{ML}, \xi) - \text{compl}(S)$

Bayesian information criterion (BIC)

Maximize:
 $score(S) = \log P(D | S, \Theta_{ML}, \xi) - \frac{1}{2} \text{compl}(S) \log N$

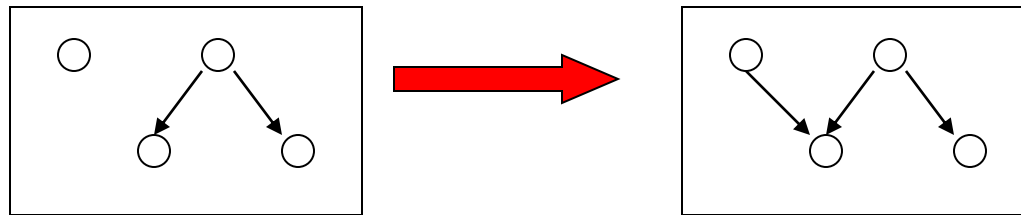
Optimizing the structure

Finding the best structure is a **combinatorial optimization** problem

- A good feature: the score **is decomposable along variables**:

$$\log P(D | S, \xi) = \sum_{i=1}^n \left\{ \sum_{j=1}^{q_i} \log \frac{\Gamma(\alpha_{ij})}{\Gamma(\alpha_{ij} + N_{ij})} + \sum_{k=1}^{r_i} \log \frac{\Gamma(\alpha_{ijk} + N_{ijk})}{\Gamma(\alpha_{ijk})} \right\}$$

Algorithm idea: Search the space of structures using local changes (additions and deletions of a link)



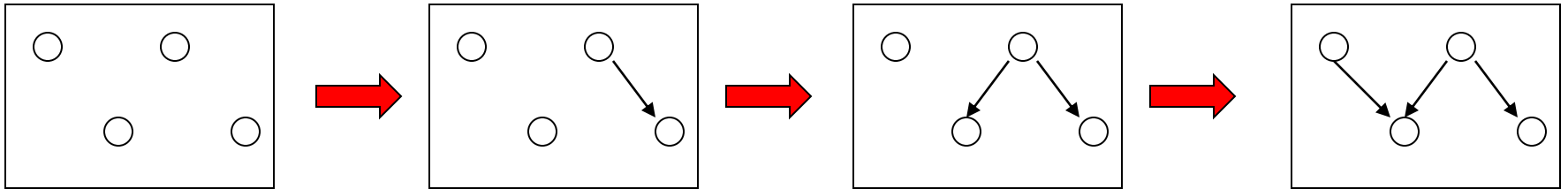
Advantage:

- we do not have to compute the whole score from scratch
- Recompute the partial score for the affected variable

Optimizing the structure. Algorithms

- **Greedy search**

- Start from the structure with no links
- Add a link that yields the best score improvement



- **Metropolis algorithm (with simulated annealing)**

- Local additions and deletions
- Avoids being trapped in “local” optimal

