

Review of Probability Distributions

CS1538: Introduction to Simulations

Some Well-Known Probability Distributions

- ▶ Bernoulli
- ▶ Binomial
- ▶ Geometric
- ▶ Negative Binomial
- ▶ Poisson
- ▶ Uniform
- ▶ Exponential
- ▶ Gamma
- ▶ Erlang
- ▶ Gaussian/Normal

Relevance to simulations:

- Need to use distributions that are appropriate for our problem
- The closer the chosen distribution matches the distribution in reality, the more accurate our model
- Why not always make a user-defined distribution specific to our problem?

Discrete Distributions

- ▶ Bernoulli
- ▶ Binomial
- ▶ Geometric
- ▶ Negative Binomial

- ▶ Poisson

The Bernoulli Distribution

- ▶ **Bernoulli Trial**: an experiment with only two possible outcomes: Success or Failure
- ▶ The probability of success is p (where $0 \leq p \leq 1$)
 - ▶ Let q be the probability of failure. *What is q in terms of p ?*
- ▶ Examples?



The Bernoulli Distribution

- ▶ Let B be a random variable over the outcome of the experiment:
 - ▶ $B = 1$ for a success; $B = 0$ for a failure
- ▶ Expectation of B
 - ▶ $E[B] = 1 \cdot p + 0 \cdot (1-p) = p$
- ▶ Variance of B
 - ▶ $\text{Var}(B) = E[B^2] - (E[B])^2 = (1^2 p + 0^2 (1-p)) - p^2 = p - p^2 = p(1-p)$



Binomial Distribution

- ▶ Experiment: Repeat Bernoulli Trial for n times
 - ▶ Good for determining the probability of getting k defective items in a batch size of n
 - ▶ Let random variable X be the number of successes
 - ▶ Note: the order doesn't matter
 - ▶ For example: suppose we toss a biased coin 3 times (with heads=success).
 - $X(\text{HHT})=2$; $X(\text{THH})=2$; $X(\text{HTH})=2$



Binomial Distribution

- ▶ The probability mass function for X :

$$p(x) = \begin{cases} \binom{n}{x} p^x q^{n-x}, & x = 0, 1, \dots, n \\ 0, & \text{otherwise} \end{cases}$$

- ▶ $E[X] = E[B_1 + B_2 + \dots B_n] = \sum E[B_i] = np$
- ▶ $\text{Var}(X) = \text{Var}(B_1 + B_2 + \dots B_n) = \sum \text{Var}(B_i) = np(1-p)$
- ▶ Note: $\binom{n}{k} = \frac{n!}{k!(n-k)!}$



Geometric Distribution

- ▶ Keep on repeating Bernoulli trials until successful
- ▶ Let r.v. X be the number of trials until the first success
- ▶ The probability mass function for X :

$$p(x) = \begin{cases} q^{x-1} p, & x = 1, 2, \dots \\ 0, & \text{otherwise} \end{cases}$$

- ▶ What is q ?



Geometric Distribution

- ▶ $E[X] = \sum xq^{(x-1)}p = p \sum xq^{(x-1)} = 1/p$
 - ▶ We know that $\int \sum xq^{(x-1)} dq = \sum q^x$ where $x = 1, 2, \dots$
 - ▶ We also know that $\sum q^y$ for $y = 0, 1, 2, \dots = 1/(1-q)$
 - ▶ So $\int \sum xq^{(x-1)} dq = 1/(1-q) - 1 = q/(1-q)$
 - ▶ $d/dq q/(1-q) = q/(1-q)^2 + 1/(1-q) = 1/(1-q)^2 = 1/p^2$
 - ▶ So $p \sum xq^{(x-1)} = d/dq \int \sum xq^{(x-1)} dq = p (1/p^2) = 1/p$
- ▶ $\text{Var}(X) = q/p^2$



Example Question

- ▶ Suppose a product has a $p\%$ chance of failure. What's the chance that it will still be working after 3 uses?



Example Question

- ▶ What's the chance that the product is still working after 7 uses, given that it works after 3 uses?



Geometric Distribution Properties

- ▶ $\Pr(X > t) = q^t$
- ▶ Memoryless
 - ▶ $\Pr(X > s+t \mid X > s) = \Pr(X > t)$

Negative Binomial Distribution

- ▶ Keep on running Bernoulli Trials until we get **k** successes.
 - ▶ Like Geometric, but now k successes instead of 1 success
- ▶ Let r.v. X be the total number of trials
- ▶ The probability mass function for X :

$$p(x) = \begin{cases} \binom{x-1}{k-1} q^{x-k} p^k, & x = k, k+1, k+2, \dots \\ 0, & \text{otherwise} \end{cases}$$

- ▶ Expectation: $E[X] = k/p$
- ▶ $\text{Var}(X) = kq/p^2$

Poisson Distribution

- ▶ Computes the probability of the number of events that may occur in a period, given the rate of occurrence in that period
- ▶ Good for modeling arrivals
- ▶ Probability mass function:

$$p(x) = \begin{cases} \frac{e^{-\alpha} \alpha^x}{x!}, & x = 0, 1, \dots \\ 0, & \text{otherwise} \end{cases}$$

- ▶ Where α is a fixed value that must be positive. It represents the average rate of the event of interest occurring.



Example

- ▶ The Prussian Cavalry/Horse study [Bortkiewicz, 1898; cf. Larsen&Marx]
 - ▶ 10 cavalry corps monitored over 20 years.
 - ▶ X = number of fatality due to horse kicks

x = # of deaths	Observed # of corps-years in which x fatality occurred	Expected # of corps-year using Poisson
0	109	
1	65	
2	22	
3	3	
4	1	
Total	200	

Poisson Distribution

- ▶ Cumulative distribution function:

$$F(x) = \sum_{i=0}^x \frac{e^{-\alpha} \alpha^i}{i!}$$

- ▶ Note that when $x \rightarrow \infty$, $F(x) \rightarrow 1$

$$\sum_{i=0}^{\infty} \frac{e^{-\alpha} \alpha^i}{i!} = e^{-\alpha} \sum_{i=0}^{\infty} \frac{\alpha^i}{i!} = e^{-\alpha} e^{\alpha} = 1$$

- ▶ $E[X] = \alpha$
- ▶ $\text{Var}(X) = \alpha$



Example

- ▶ Suppose that I see students at the rate of five per office hours. What is the chance that I'll see 4 students at office hours today?
- ▶ What is the chance that I'll see less than 3 students at office hours today?
- ▶ What is the chance that I'll see at least 3 students at office hours today?

Poisson: Relationship to Binomial

- ▶ **Theorem:** Let α be a fixed number and n be an arbitrary positive integer. For each nonnegative integer x ,

$$\lim_{n \rightarrow \infty} \binom{n}{x} p^x (1-p)^{n-x} = \frac{e^{-\alpha} \alpha^x}{x!}$$

where $p = \alpha/n$

- ▶ This makes it easier to calculate the binomial distribution, especially for large n 's

Poisson Arrival Process

- ▶ Recall α represents average rate of occurrence
 - ▶ How do we represent time more explicitly?
- ▶ Let $N(t)$ be a random variable that represents the number (positive integer) of events that occurred in time $[0, t]$
 - ▶ We want to know $\Pr(N(t) = n)$
 - ▶ If we guarantee a few properties, we can use the Poisson Distribution

Poisson Arrival Process

► Assumptions:

- Arrivals occur one at a time
- Arrival rate (λ) does not change over time
 - We'll see how to relax this
- The number of arrivals in given period are independent of each other

► We can now rewrite the Poisson Distribution:

► $\alpha = \lambda t$

$$P[N(t) = n] = \begin{cases} \frac{e^{-\lambda t} (\lambda t)^n}{n!}, & n = 0, 1, \dots \\ 0, & \text{otherwise} \end{cases}$$

► How does this affect the expectation and variance?

Properties of the Poisson Arrival Process

▶ Random Splitting

- ▶ Consider a Poisson Process $N(t)$ with rate λ . Assume that arrivals can be divided into two groups, A and B with probability p and $(1-p)$, respectively
- ▶ N_A is a Poisson Process with rate λp and N_B is a Poisson Process with rate $\lambda(1-p)$
- ▶ $N(t) = N_A(t) + N_B(t)$



Properties of the Poisson Arrival Process

▶ Pooled Process

- ▶ Consider two Poisson Processes $N_1(t)$ and $N_2(t)$, with rates λ_1 and λ_2
- ▶ The sum of the two processes is also a Poisson Process; it has a rate of $\lambda_1 + \lambda_2$
- ▶ Pooling can be used in situations where multiple arrival processes feed a single queue



Discrete Probability Summary

- ▶ Binomial Distribution
- ▶ Geometric Distribution
- ▶ Negative Binomial Distribution

These describe experiments based on repeated Bernoulli Trials

- ▶ Poisson Distribution

Doesn't have an easy to describe underlying structure, but seems to be a good fit for many real data sets



Continuous Random Variables

- ▶ Random variable X is **continuous** if its *range space* is an interval or a collection of intervals
- ▶ There exists a non-negative function $f(x)$, called the **probability density function**, such that for any set of real numbers,
 - ▶ $f(x) \geq 0$ for all x in the range space
 - ▶ $\int_{\text{rangespace}} f(x) dx = 1$ (i.e., the total area under $f(x)$ is 1)
 - ▶ $f(x) = 0$ for all x not in the range space

Note that $f(x)$ does **not give the probability of $X = x$**

- ▶ Unlike the pmf for discrete random variables



Continuous Random Variables

- ▶ The probability that X lies in a given interval $[a,b]$ is

$$P(a \leq X \leq b) = \int_a^b f(x)dx$$

- ▶ aka "area under the curve"
- ▶ Note that for continuous random variables,
 $\Pr(X = x) = 0$ for any x
- ▶ Consider the probability of x within a (very small) range
- ▶ The **cumulative distribution function** (cdf), $F(x)$ is now the integral from $-\infty$ to x or

$$F(x) = \int_{-\infty}^x f(t)dt$$

- This gives us the probability up to x



Continuous Random Variables

- ▶ **Expected Value** for a continuous random variable
 - ▶ Similar to the discrete case, except that we integrate instead of summing

$$E[X] = \int_{-\infty}^{\infty} xf(x)dx$$

- ▶ **Variance**: same formulation as its discrete counterpart (though calculating $E[X^2]$ will involve integrals again).

$$\text{Var}(X) = E[X^2] - (E[X])^2$$



Uniform Distribution over range [a,b]

- ▶ Probability density function:

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{if } a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

- ▶ Cumulative distribution function:

$$F(x) = \int_a^x f(y)dy = \int_a^x \frac{1}{b-a} dy = \frac{x-a}{b-a} \quad \text{if } a \leq x \leq b$$

- ▶ What about $F(x)$ when $x < a$ or $x > b$?



Uniform Distribution over range [a,b]

► **Expectation** $E(X) = \int_a^b \frac{x}{b-a} dx = \left. \frac{x^2}{2(b-a)} \right|_a^b = \frac{b^2 - a^2}{2(b-a)} =$

$$\frac{(b+a)(b-a)}{2(b-a)} = \frac{b+a}{2}$$

- What does the expected value of a discrete uniform distribution look like?

► **Variance** $Var(X) = E[X^2] - (E[X])^2 = \int_a^b \frac{x^2}{b-a} dx - (E[X])^2$

$$= \left. \frac{x^3}{3(b-a)} \right|_a^b - \frac{(b+a)^2}{4} = \frac{b^3 - a^3}{3(b-a)} - \frac{(b+a)^2}{4}$$
$$= \frac{4(b^3 - a^3) - 3(b+a)^2(b-a)}{12(b-a)} = \frac{(b-a)^2}{12}$$



Normal Distribution

- ▶ Probability density function:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp \left[-\frac{1}{2} \left(\frac{x-\mu}{\sigma} \right)^2 \right], \quad -\infty < x < \infty$$

where we supply the mean and variance:

- ▶ μ : mean
 - ▶ σ : square-root of variance
 - ▶ The normal distribution is also denoted as: $\mathbf{N}(\mu, \sigma^2)$
- ▶ Some of its properties:

$$\lim_{x \rightarrow -\infty} f(x) = 0 = \lim_{x \rightarrow \infty} f(x)$$

$$f(\mu + x) = f(\mu - x)$$

$$\max(f(x)) = f(\mu)$$



Normal Distribution

$$\blacktriangleright F(x) = P(X \leq x) = \int_{-\infty}^x \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{t-\mu}{\sigma}\right)^2\right] dt$$

- ▶ Doesn't have a closed form.
- ▶ Can use numerical methods, but want to avoid evaluating integrals for each pair (μ, σ^2)

- ▶ Transform to **standard normal distribution**

- ▶ Let $z = (t-\mu)/\sigma$ then we can rewrite the above as:

- ▶ $F(x) = P(X \leq x) = P(Z \leq (x-\mu)/\sigma)$

$$= \int_{-\infty}^{(x-\mu)/\sigma} \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{1}{2}(z)^2\right] dz$$



CDF table for $N(0,1)$

- ▶ There are a few variants. On the website:

<i>Z</i>	.00	.01	.02	.03
-3.9	.00005	.00005	.00004	.0000
-3.8	.00007	.00007	.00007	.0000
-3.7	.00011	.00010	.00010	.0001
-3.6	.00016	.00015	.00015	.0001
-3.5	.00023	.00022	.00022	.0002
-3.4	.00034	.00032	.00031	.0003
-3.3	.00048	.00047	.00045	.0004
-3.2	.00069	.00066	.00064	.0006

- ▶ A typical table from a stats book lets you specify z to 2 significant digits: ($z = \text{column}_1 + \text{row}_1$)



Example (from textbook 5.21)

- ▶ Suppose we have a normal distribution such that:

X is a r.v. from $N(50, 9)$

- ▶ What is the chance that $X \leq 56$?
 - ▶ We wish to compute $F(56)$ (i.e., $P(X \leq 56)$).



Exponential Distribution

- ▶ **Models**
 - ▶ Interarrival times
 - ▶ Service times
 - ▶ Lifetime of a component that fails instantaneously
- ▶ **Parameter λ indicates rate**

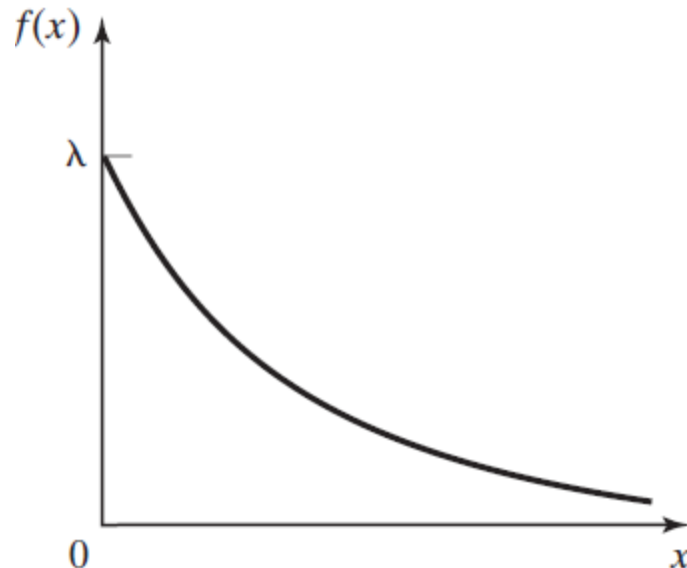


Exponential Distribution

- ▶ Probability density function:

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

- ▶ where $\lambda > 0$ is a parameter that we supply
- ▶ Since the exponent is negative, the pdf will decrease as x increases

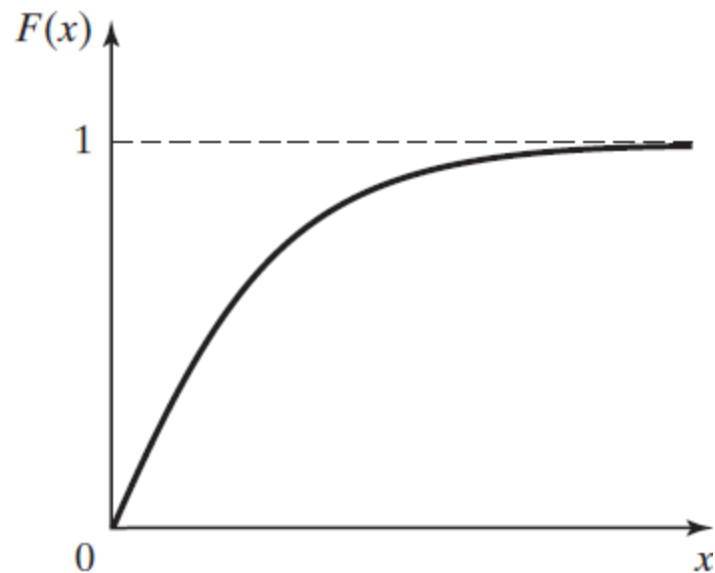


Exponential Distribution

$$F(x) = \begin{cases} 0, & x < 0 \\ \int_0^x \lambda e^{-\lambda t} dt = 1 - e^{-\lambda x}, & x \geq 0 \end{cases}$$

$$E[X] = \frac{1}{\lambda}$$

$$\text{Var}(X) = \frac{1}{\lambda^2}$$



Exponential Distribution: λ

- ▶ λ represents a rate: number of occurrences per time unit
- ▶ x : amt. of time it took for some occurrence to take place
- ▶ $F(x)$: prob. that the event happened during interval $[0, x]$
- ▶ $1 - F(x)$: prob. that the event doesn't happen until after x

Example 5.17

Suppose lifespan of a type of lamp is exponentially distributed with failure rate $\lambda = 1/3000$ hrs. What's the chance a particular lamp will beat the average?

- ▶ Has a relationship to the **Poisson Arrival Process**



Exponential Distribution

- ▶ Like the geometric distribution, the exponential distribution is **memoryless**
 - ▶ $P(X > s+t \mid X > s) = P(X > t)$
 - ▶ The proof is similar to how we showed that the geometric distribution is memoryless



Relationship between Poisson Distribution and Exponential Distribution (1 / 3)

- ▶ Events occur at a rate of λ
- ▶ Poisson: chance of n events taking place is given by:

$$\Pr[N(t)=n] = e^{-\lambda t} (\lambda t)^n / n!$$

- ▶ Let E be the first occurrence of an event since the start of the clock at 0; let t represent the time when E happens
- ▶ Want to find: distribution for t
- ▶ Look for CDF $F(t)$ – the probability that E happened some time within the interval $[0,t]$
 - ▶ Then we can take the derivative of $F(t)$ to find $f(t)$
 - ▶ Want to show: $F(t)$ fits Exponential



Relationship between Poisson Distribution and Exponential Distribution (2/3)

- ▶ $F(t)$ = The chance that E occurs in some interval $[0,t]$
= $1 - \Pr(\text{E occurred after } t) = 1 - \Pr(\text{nothing took place in } [0,t])$
- ▶ *Nothing took place in $[0,t]$* relates counting discrete event occurrences with measuring continuous time duration
 - ▶ $\Pr(\text{nothing took place in } [0,t]) = \Pr[N(t)=0]$
 - ▶ $F(t) = 1 - \Pr(\text{E occurred after } t) = 1 - \Pr[N(t) = 0] = 1 - e^{-\lambda t}$
which is the CDF for exponential distribution
- ▶ So the duration between the 0th and 1st event of a Poisson arrival process follows an exponential distribution



Relationship between Poisson Distribution and Exponential Distribution (3/3)

- ▶ More generally, for any two events E_i and E_j that follows the Poisson arrival process, the duration between them also follows the exponential distribution
- ▶ Let i be the time when E_i occurs and j be the time when E_j occurs
- ▶ We can use the same analysis as before (just align i with 0 and j with t in the previous case)
 - ▶ This is because Poisson arrival process assumes constant rate
 - ▶ Also makes sense – exponential distribution is memoryless



Exercises with Exponential

- ▶ Rate: 0.5 arrivals per minute
- ▶ $F(t) = 1 - e^{-0.5t}$ = probability of time elapse between arrivals
 - ▶ What's the probability that the next customer shows up before 30 seconds have passed?
 - ▶ ... within 1 minute?
 - ▶ ... within 2 minutes?
- ▶ What's the chance that a customer shows up between 5 and 7 minutes?



Relationship between Gamma Distribution and Poisson Distribution

- ▶ Suppose some event occurs over time interval of length x at the average rate of λ per unit time. How long would it take for r events to happen?
 - ▶ Divide up x into small independent subintervals such that the chance of more than one event happening in the subinterval is negligibly small
 - ▶ Let W be a r.v. counting the # of occurrences of the event in the total duration $[0,x]$. W is a Poisson r.v. w/ params λx
 - ▶ $F_X(x) = \Pr(X \leq x) = 1 - \Pr(X > x)$
$$= 1 - F_W(r-1) = 1 - \sum_{k=0}^{r-1} e^{-\lambda x} \frac{(\lambda x)^k}{k!}$$
 - ▶ $f_X(x) = F'_X(x) = \frac{\lambda^r}{(r-1)!} x^{r-1} e^{-\lambda x}$



The Gamma function:

- ▶ If r can be non-integer, then we need to replace $(r-1)!$ with a continuous function of r .
- ▶ We'll call this function *Gamma* of r : $\Gamma(r)$
 - ▶ For any real number $r > 0$, the gamma function $\Gamma(r)$ is:

$$\Gamma(\beta) = \int_0^{\infty} x^{\beta-1} e^{-x} dx$$

- ▶ $\Gamma(1) = 1$
- ▶ $\Gamma(1/2) = \text{sqrt}(\pi)$
- ▶ $\Gamma(r+1) = r \Gamma(r)$ for any positive real r
- ▶ $\Gamma(r+1) = r!$ if r is a nonnegative integer
- ▶ $\binom{n+r-1}{n} = \frac{\Gamma(n+r)}{\Gamma(n+1)\Gamma(r)}$
- ▶ $\frac{\Gamma(r)\Gamma(s)}{\Gamma(r+s)} = \int_0^1 u^{r-1}(1-u)^{s-1} du$



Gamma Distribution (general form)

- ▶ Useful for when waiting times between events is relevant (e.g. waiting time between Poisson events)
- ▶ Let X be a random variable such that

$$f_X(x) = \begin{cases} \frac{\beta\theta}{\Gamma(\beta)} (\beta\theta x)^{\beta-1} e^{-\beta\theta x}, & x > 0, \beta > 0 \\ 0, & \text{otherwise} \end{cases}$$

- ▶ β is referred to as a **shape** parameter
- ▶ θ is referred to as a **scale** parameter
- ▶ $E[X] = \int_0^{\infty} x f(x) dx = \frac{1}{\theta}$
- ▶ $\text{Var}(X) = \frac{1}{\beta\theta^2}$
- ▶ CDF: $F_X(x) = \int_0^x f(t) dt = 1 - \int_x^{\infty} f(t) dt$



Erlang Distribution

- ▶ When β is an arbitrary positive integer, the Gamma Distribution is called the **Erlang Distribution of order k** ($k = \beta$)

- ▶ Can simplify the CDF to:

$$F(X) = \begin{cases} 1 - \sum_{i=0}^{k-1} \frac{e^{-k\theta x} (k\theta x)^i}{i!} & x > 0 \\ 0 & x \leq 0 \end{cases}$$

- ▶ This is the sum of Poisson terms with mean $\alpha = k\theta x$

- ▶ **k** is the number of events
- ▶ **θ** is the rate of the collection of events
- ▶ **$\lambda = k\theta$** is the rate of one event (events per unit time)



Erlang Example

- ▶ Suppose a node in a network does not transmit until it has accumulated 5 messages in its buffer. Suppose messages arrive independently and are exponentially distributed with a mean of 100 ms between messages.
- ▶ Suppose a transmission was just made; what's the probability that more than 552 ms will pass before the next transmission?