

CS 3750 Advanced Machine Learning

Latent Variable Generative Models II

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Based on slides of Professor Milos Hauskrecht

Outline

- Latent Variable Generative Models
- Cooperative Vector Quantizer Model
 - Model Formulation
 - Expectation Maximization (EM)
 - Variational Approximation
- Noisy-OR Component Analyzer
 - Model Formulation
 - Variational EM for NOCA
- References

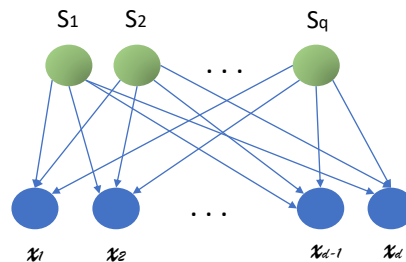
Latent Variable Generative Models

- Generative Models: Unsupervised learning models that study the underlying structure (e.g. interesting patterns) and causal structures of data to generate data like it.
- Latent (hidden) variables are random variables that are hard to observe. (ex. Length is measured, but intelligence is not), and is assumed to affect the response variable.
- The idea: introduce an unobserved latent variable, S , and use it to generate a traceable, less complex distribution.

$$\begin{array}{ccc}
 p(x) & \longrightarrow & p(x, s) = p(x | s) p(s) \\
 \text{Complex Distribution} & & \text{Simpler Distribution}
 \end{array}$$

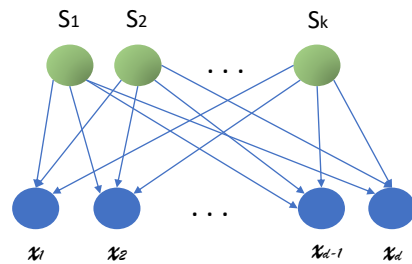
Latent Variable Generative Models

- Assumption: Observable variables are independent given latent variables.



Cooperative Vector Quantizer (CVQ)

- Latent variables (s): Binary vars with Dimensionality k
- Observed variables (x): real valued vars Dimensionality d



CVQ – Model Description

- Model

$$\mathbf{x} = \sum_{k=1}^K s_k \mathbf{w}_k + \epsilon$$

- Latent variables s_i

- \sim Bernoulli distribution parameter: π_i
- $P(s_i | \pi_i) = \pi_i^{s_i} (1 - \pi_i)^{1-s_i}$
- w_k is the weight output by source s_k

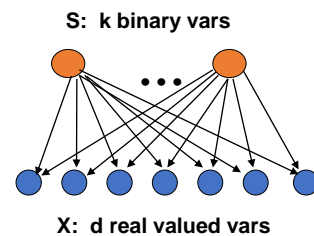
- Observable variables x

- \sim Normal distributions parameters: \mathbf{W}, Σ
- $P(x | s) = N(\mathbf{W}s, \Sigma)$,
- we assume $\Sigma = \sigma I$

$$\mathbf{W} = \begin{pmatrix} w_{11} & w_{12} & \dots & w_{1k} \\ w_{21} & & & \\ & \dots & & \\ w_{d1} & \dots & \dots & w_{dk} \end{pmatrix}$$

- Joint for one instance of s and x

$$p(x, s | \theta) = 2^{-d/2} \sigma^{-d/2} \exp\left\{-\frac{1}{2\sigma^2} (x - \mathbf{W}s)^T (x - \mathbf{W}s)\right\} \prod_{i=1}^k \pi_i^{s_i} (1 - \pi_i)^{1-s_i}$$



CVQ – Model Description

- **Objective:** to learn parameters of the model: W, π, σ
- If both x and s are observable,
 - Use loglikelihood:

$$\sum_{n=1}^N \log P(x^{(n)}, s^{(n)} | \Theta) = \sum_{n=1}^N -d \log \sigma - \frac{1}{2\sigma^2} (x^{(n)} - Ws^{(n)})^T (x^{(n)} - Ws^{(n)}) + \sum_{i=1}^k s_i^{(n)} \log \pi_i (1 - s_i^{(n)}) \log(1 - \pi_i) + c$$

- Solution is nice and easy

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CVQ – Model Description

- **Objective:** to learn parameters of the model: W, π, σ
- If only x are observable
 - Log likelihood of data:

$$\log P(D | \Theta) = \sum_{n=1}^N \log P(x^{(n)} | \Theta) = \sum_{n=1}^N \log \sum_{\{s^{(n)}\}} P(x^{(n)}, s^{(n)} | \Theta)$$

- Solution is hard, we can no longer benefit from the decomposition.
- Use Expectation Maximization (EM).

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Expectation Maximization (EM)

- Let H be a set of all variables with hidden or missing values
 - $P(H, D | \Theta, \xi) = P(H | D, \Theta, \xi)P(D | \Theta, \xi)$
 - $\log P(H, D | \Theta, \xi) = \log P(H | D, \Theta, \xi) + \log P(D | \Theta, \xi)$
 - $\log P(D | \Theta, \xi) = \log P(H, D | \Theta, \xi) - \log P(H | D, \Theta, \xi)$
- Average both sides with $P(H | D, \Theta', \xi)$ for Θ'
 - $E_{H|D,\Theta'} \log P(D | \Theta, \xi) = E_{H|D,\Theta'} \log P(H, D | \Theta, \xi) - E_{H|D,\Theta'} \log P(H | D, \Theta, \xi)$
 - $\log P(D | \Theta, \xi) = F(\Theta | \Theta') = E(\Theta | \Theta') + H(\Theta | \Theta')$
- EM uses the true posterior. $P(H | D, \Theta', \xi)$

Log-likelihood
of data

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Expectation Maximization (EM)

- **General EM Algorithm:**
 - Initialize parameters Θ
 - Set $\Theta' = \Theta$
- **Expectation step**
 - $E(\Theta | \Theta') = \langle \log P(H, D | \Theta, \xi) \rangle_{P(H | D, \Theta')}$
- **Maximization step**
 - $\Theta = \operatorname{argmax} E(\Theta | \Theta')$
 - Repeat until no or small improvement in Θ ($\Theta = \Theta'$)
- **Problem**
 - $P(H | D, \Theta') = \prod_{n=1}^N P(x^{(n)}, s^{(n)} | \Theta')$
 - Each data point requires us to calculate 2^k probabilities
 - If k is large, then this is a bottleneck

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Variational Approximation

- An alternative method to approximate inference based on stochastic sampling.
- Let H be a set of all variables with hidden or missing values
 - $\log P(D|\Theta, \xi) = \log P(H, D|\Theta, \xi) - \log P(H|D, \Theta, \xi)$
- Average both sides using a distribution $Q(H|\lambda)$ [*surrogate posterior*]

$$E_{H|\lambda} \log P(D|\Theta, \xi) = E_{H|\lambda} \log P(H, D|\Theta, \xi) - E_{H|\lambda} \log Q(H|\lambda) \\ + E_{H|\lambda} \log Q(H|\lambda) - E_{H|\lambda} \log P(H|D, \Theta, \xi)$$

$$\log P(D|\Theta, \xi) = F(Q, \Theta) + KL(Q, P)$$

$$F(Q, \Theta) = \sum_{\{H\}} Q(H|\lambda) \log P(H, D|\Theta, \xi) - \sum_{\{H\}} Q(H|\lambda) \log Q(H|\lambda)$$

$$KL(Q, P) = \sum_{\{H\}} Q(H|\lambda) [\log Q(H|\lambda) - \log P(H|D, \Theta)]$$

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Variational Approximation

$$\log P(D|\Theta, \xi) = F(Q, \Theta) + KL(Q, P)$$

$$F(Q, \Theta) = \sum_{\{H\}} Q(H|\lambda) \log P(H, D|\Theta, \xi) - \sum_{\{H\}} Q(H|\lambda) \log Q(H|\lambda)$$

$$KL(Q, P) = \sum_{\{H\}} Q(H|\lambda) [\log Q(H|\lambda) - \log P(H|D, \Theta)]$$

- **Approximation:** maximize $F(Q, \Theta)$
- **Parameters:** Θ, λ
- Maximization of F pushes up the lower bound on the log-likelihood

$$\log P(D|\Theta, \xi) \geq F(Q, \Theta).$$

Kullback-Leibler (KL) divergence

- A method to measure the difference between two probability distributions over the same variable x
 - $KL(P \parallel Q)$
 - Where the “ \parallel ” operator indicates “*divergence*” or P ’s divergence from Q
- Entropy: the average amount of information for a probability distribution
 - $H(P) = E_P[I_P(X)] = - \sum_{i=1}^n P(i) \log(P(i))$
 - $KL(P \parallel Q) = H(P, Q) - H(P) = - \sum_{i=1}^n P(i) \log(Q(i)) + \sum_{i=1}^n P(i) \log(P(i)) = \sum_{i=1}^n P(i) \log\left(\frac{P(i)}{Q(i)}\right)$
- If we have some theoretic minimal distribution P , we want to try to find an approximation Q that tries to get as close as possible by minimizing the KL divergence

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Variational EM

- To use Variational EM, we hope if we choose $Q(H | \lambda)$ well, the optimization of both λ and Θ will become easy.
- A well-behaved choice for $Q(H | \lambda)$ is the *mean field approximation*.
- Let H – be a set of all variables with hidden or missing values:
 - **E-step:** *Compute expectation over hidden variables*
 - Optimize: $F(Q, \Theta)$ with respect to λ while keeping Θ fixed.
 - **M-step:** *Maximize expected loglikelihood*
 - Optimize: $F(Q, \Theta)$ with respect to Θ while keeping λ s fixed.

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Mean Field Approximation

- To find the distribution Q , we use Mean Field Approximation

- **Assumption:**

- $Q(H|\lambda)$ is the *mean field approximation*
- Variables in the $Q(H)$ distribution are *independent* variables H_i
- Q is completely *factorized*

$$Q(H|\lambda) = \prod Q_i(H_i|\lambda_i)$$

- For our CVQ model

- Hidden variables are binary sources

$$Q(H|\lambda) = \prod_{n=1 \dots N} Q(s^{(n)}|\lambda^n)$$

$$Q(s^{(n)}|\lambda^n) = \prod_{i=1 \dots k} Q(s_i^{(n)}|\lambda_i^{(n)})$$

$$Q(s_i^{(n)}|\lambda_i^{(n)}) = \lambda_i^{(n)s_i^{(n)}} (1 - \lambda_i^{(n)})^{1-s_i^{(n)}}$$

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Mean Field Approximation

- Functional F for the mean field:

$$F(Q, \theta) = \sum_{\{H\}} Q(H|\lambda) \log P(H, D|\theta, \xi) - \sum_{\{H\}} Q(H|\lambda) \log Q(H|\lambda)$$

- Assume just one data point \mathbf{x} and corresponding \mathbf{s} :

$$F(Q, \theta) = \sum_{n=1}^N \langle \log P((x^{(n)}, s^{(n)}|\theta)) \rangle_{Q(s^{(n)}|\lambda^{(n)})} - \langle \log Q(s^{(n)}|\lambda^{(n)}) \rangle_{Q(s^{(n)}|\lambda^{(n)})}$$

$$= \langle -d \log \sigma - \frac{1}{2\sigma^2} (\mathbf{x} - \mathbf{W}\mathbf{s})^T (\mathbf{x} - \mathbf{W}\mathbf{s}) \rangle_{Q(s|\lambda)} \quad (1)$$

$$+ \langle \sum_{i=1}^k s_i \log \pi_i + (1 - s_i) \log(1 - \pi_i) \rangle_{Q(s|\lambda)} \quad (2)$$

$$- \langle \sum_{i=1}^k s_i \log \lambda_i + (1 - s_i) \log(1 - \lambda_i) \rangle_{Q(s|\lambda)} \quad (3)$$

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Mean Field Approximation

- Functional F. Part (1)

$$\begin{aligned}
 & \langle -d \log \sigma - \frac{1}{2\sigma^2} (\mathbf{x} - \mathbf{W}\mathbf{s})^T (\mathbf{x} - \mathbf{W}\mathbf{s}) \rangle_{Q(s|\lambda)} \\
 &= \langle -d \log \sigma - \frac{1}{2\sigma^2} (\mathbf{x} - \sum_{i=1}^k s_i \mathbf{w}_i)^T (\mathbf{x} - \sum_{i=1}^k s_i \mathbf{w}_i) \rangle_{Q(s|\lambda)} \\
 &= \langle -d \log \sigma - \frac{1}{2\sigma^2} \left[\mathbf{x}^T \mathbf{x} - 2 \sum_{i=1}^k (s_i \mathbf{w}_i)^T \mathbf{x} + \sum_{i=1}^k \sum_{j=1}^k s_i s_j \mathbf{w}_i^T \mathbf{w}_j \right] \rangle_{Q(s|\lambda)} \\
 &= -d \log \sigma - \frac{1}{2\sigma^2} \left[\mathbf{x}^T \mathbf{x} - 2 \sum_{i=1}^k \langle s_i \rangle_{Q(s_i|\lambda_i)} \mathbf{w}_i^T \mathbf{x} + \sum_{i=1}^k \sum_{j=1}^k \langle s_i s_j \rangle_{Q(s|\lambda)} \mathbf{w}_i^T \mathbf{w}_j \right] \\
 \\
 & \langle s_i \rangle_{Q(s_i|\lambda_i)} = \lambda_i \qquad \langle s_i s_j \rangle_{Q(s|\lambda)} = \lambda_i \lambda_j + \delta_{ij} (\lambda_i - \lambda_i^2)
 \end{aligned}$$

Mean Field Approximation

- Functional F. Part (2)

$$\begin{aligned}
 \langle \sum_{i=1}^k s_i \log \pi_i + (1 - s_i) \log(1 - \pi_i) \rangle_{Q(s|\lambda)} &= \sum_{i=1}^k \langle s_i \rangle_{Q(s_i|\lambda_i)} \log \pi_i + (1 - \langle s_i \rangle_{Q(s_i|\lambda_i)}) \log(1 - \pi_i) \\
 &= \sum_{i=1}^k \lambda_i \log \pi_i + (1 - \lambda_i) \log(1 - \pi_i)
 \end{aligned}$$

- Functional F. part (3)

$$\langle \sum_{i=1}^k s_i \log \lambda_i + (1 - s_i) \log(1 - \lambda_i) \rangle_{Q(s|\lambda)} = \sum_{i=1}^k \lambda_i \log \lambda_i + (1 - \lambda_i) \log(1 - \lambda_i)$$

Mean Field Approximation

Functional F:

$$\begin{aligned}
 &= -d \log \sigma - \frac{1}{2\sigma^2} \left[\mathbf{x}^T \mathbf{x} - 2 \sum_{i=1}^k \langle s_i \rangle_{Q(s_i | \lambda_i)} \mathbf{w}_i \right] \mathbf{x} + \sum_{i=1}^k \sum_{j=1}^k \langle s_i s_j \rangle_{Q(s_i | \lambda)} \mathbf{w}_i^T \mathbf{w}_j \\
 &+ \sum_{i=1}^k \lambda_i \log \pi_i + (1 - \lambda_i) \log(1 - \pi_i) \\
 &+ \sum_{i=1}^k \lambda_i \log \lambda_i + (1 - \lambda_i) \log(1 - \lambda_i)
 \end{aligned}$$

Parameters: W, π, σ

Mean field parameters: λ

Mean Field Approximation

Functional F (for all data points):

$$\begin{aligned}
 F(Q, \theta) &= \sum_{n=1}^N \langle \log P(\mathbf{x}^{(n)}, \mathbf{s}^{(n)} | \theta) \rangle_{Q(\mathbf{s}^{(n)} | \lambda^{(n)})} - \langle \log Q(\mathbf{s}^{(n)} | \lambda^{(n)}) \rangle_{Q(\mathbf{s}^{(n)} | \lambda^{(n)})} \\
 &= -d \log \sigma - \frac{1}{2\sigma^2} \left[\mathbf{x}^{(n)T} \mathbf{x}^{(n)} - 2 \sum_{i=1}^k \lambda_i^{(n)} \mathbf{w}_i \right] \mathbf{x}^{(n)} + \sum_{i=1}^k \sum_{j=1}^k \left[\lambda_i^{(n)} \lambda_j^{(n)} + \delta_{ij} (\lambda_i^{(n)} - \lambda_i^{(n)2}) \right] \mathbf{w}_i^T \mathbf{w}_j \\
 &+ \sum_{i=1}^k \lambda_i^{(n)} \log \pi_i + (1 - \lambda_i^{(n)}) \log(1 - \pi_i) \\
 &+ \sum_{i=1}^k \lambda_i^{(n)} \log \lambda_i^{(n)} + (1 - \lambda_i^{(n)}) \log(1 - \lambda_i^{(n)})
 \end{aligned}$$

Parameters: W, π, σ

Mean field parameters: $\lambda = \lambda^1, \lambda^2, \dots, \lambda^n$

Variational EM

- E-step

- Optimize $F(Q, \Theta)$ with respect to λ while keeping Θ fixed

$$\frac{\partial}{\partial \lambda_u} F = \frac{1}{\sigma^2} (\mathbf{x} - \sum_{j \neq u} \lambda_j \mathbf{w}_j)^T \mathbf{w}_u - \frac{1}{2\sigma^2} \mathbf{w}_u^T \mathbf{w}_u + \log \frac{\pi_u}{1 - \pi_u} - \log \frac{\lambda_u}{1 - \lambda_u}$$

$$\text{Set } \frac{\partial}{\partial \lambda_u} F = 0$$

$$\lambda_u = \text{g}\left(\frac{1}{\sigma^2} (\mathbf{x} - \sum_{j \neq u} \lambda_j \mathbf{w}_j)^T \mathbf{w}_u - \frac{1}{2\sigma^2} \mathbf{w}_u^T \mathbf{w}_u + \log \frac{\pi_u}{1 - \pi_u}\right), \quad \text{g}(x) = \frac{1}{1 + e^{-x}}$$

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Variational EM

- M-step

- Optimize $F(Q, \Theta)$ with respect to Θ while keeping λ s

Start with π :

For N data points

$$\frac{\partial}{\partial \pi_u} F = \sum_{n=1}^N \lambda_u^n \log \frac{1}{\pi_u} - (1 - \lambda_u^n) \log \frac{1}{(1 - \pi_u)}$$

$$\text{Set } \frac{\partial}{\partial \pi_u} F = 0,$$

$$\pi_u = \frac{\sum_{n=1}^N \lambda_u^{(n)}}{N} \quad (\text{closed form solution})$$

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Variational EM

And for parameter w :

$$\frac{\partial}{\partial w_{uv}} F = \sum_{n=1}^N -\frac{1}{2\sigma^2} \left[\lambda_v^{(n)} x_u^{(n)} + 2 \sum_{j \neq v} \lambda_v^{(n)} \lambda_j^{(n)} w_{uj} + 2 \lambda_v^{(n)} w_{uv} \right] = 0$$

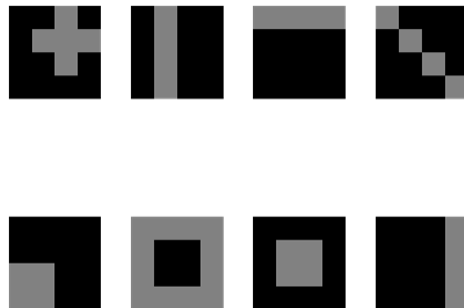
$$\mathbf{W} = \begin{pmatrix} w_{11} & w_{12} & \dots & w_{1k} \\ w_{21} & & & \\ & \dots & & \\ w_{d1} & \dots & \dots & w_{dk} \end{pmatrix} \quad \mathbf{W} = (\mathbf{w}_1 \ \mathbf{w}_2 \ \dots \ \mathbf{w}_k)$$

- For each variable v :

The equations define a set of k linear equations that can be solved

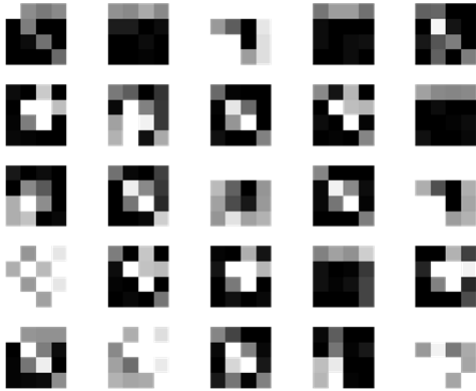
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Image Separation Experiment



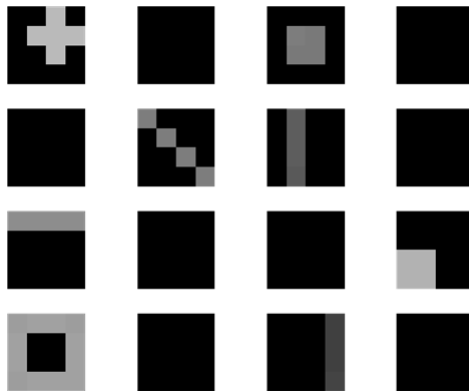
Source images associated with latent variables

Mixed images



- Images generated by the model.
- Some of the images are noise.
- Generating enough samples; the model can retrieve the original source.

Recovered sources



Modeling High-Dimensional Data

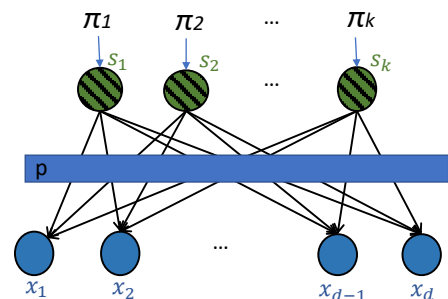
- Definition: Number of dimensions are high that makes calculations extremely difficult (number of features exceed number of observations).
- Examples of domains with High-Dimensional Data:
 - Sensor Networks
 - Document Repositories.
- Typically, variables are dependent.
- How to model dependencies?
 - Full model (intractable, overfitting)
 - All-independent (unrealistic)
 - Middle-of-the-road approaches
 - Captures dependencies in an efficient way (representation, reasoning, learning).

Noisy-OR Component Analyzer

- Objective: Capture dependencies via latent factors and combinations.
- The dependencies between observables are represented using a smaller number of *hidden binary* factors.
- NOCA model has binary nodes:
 - k parameters for each observed node, p_{1j}, \dots, p_{kj}
 - p_{ij} is interpreted as “strength of influence” of S_i on observable variable X_j

$$P(X_j = 0 | \mathbf{s}) = \prod_{i=1}^K (1 - p_{ij})^{s_i}$$

$$P(X_j = 1 | \mathbf{s}) = 1 - P(X_j = 0 | \mathbf{s}) = 1 - \prod_{i=1}^K (1 - p_{ij})^{s_i}$$



Noisy-OR Component Analyzer

- A generalization of the logical OR

Assumptions:

- All possible causes U_i for an event X are modeled using nodes (random variables) and their values, with T (or 1) reflecting the presence of the cause, and F (or 0) its absence
- If one needs to represent unknown causes one can add a leak node
- **Parameters:** For each cause U_i define an (independent) probability q_i that represents the probability with which the cause does not lead to $X = T$ (or 1), or in other words, it represents the probability that the positive value of variable X is inhibited when U_i is present

$$p(x = 1 | U_1, \dots, U_j, \neg U_{j+1}, \dots, \neg U_k) = 1 - \prod_{i=1}^j q_i$$

$$p(x = 0 | U_1, \dots, U_j, \neg U_{j+1}, \dots, \neg U_k) = \prod_{i=1}^j q_i$$

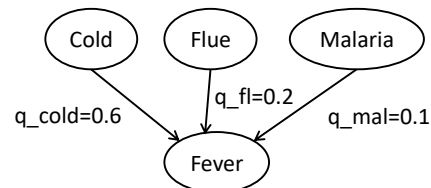
Note: The negated causes $\neg U_i$ (reflecting the absence of the cause) do not have any influence on X . Why?

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Noisy-OR Example

$$\mu(x = 1 | U_1, \dots, U_j, \neg U_{j+1}, \dots, \neg U_k) = 1 - \prod_{i=1}^j q_i$$

$$\mu(x = 0 | U_1, \dots, U_j, \neg U_{j+1}, \dots, \neg U_k) = \prod_{i=1}^j q_i$$



| Cold | Flu | Malaria | $\mu(\text{Fever})$ | $\mu(\neg \text{Fever})$ |
|------|-----|---------|---------------------|----------------------------|
| F | F | F | 0 | 1 |
| F | F | T | 0.9 | 0.1 |
| F | T | F | 0.8 | 0.2 |
| F | T | T | 0.98 | 0.02 = 0.2 × 0.1 |
| T | F | F | 0.4 | 0.6 |
| T | F | T | 0.94 | 0.06 = 0.6 × 0.1 |
| T | T | F | 0.88 | 0.12 = 0.6 × 0.2 |
| T | T | T | 0.988 | 0.012 = 0.6 × 0.2 × 0.1 |

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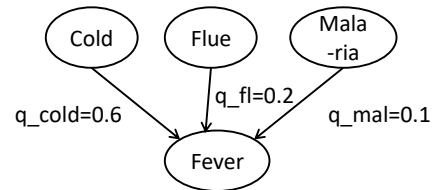
Noisy-OR parameter reduction

- Please note that in general the number of entries defining the CPT (conditional probability table) grows exponentially with the number of parents;
 - for q binary parents the number is : 2^q
- For the noisy-or CPT the number of parameters is $q + 1$

Example:

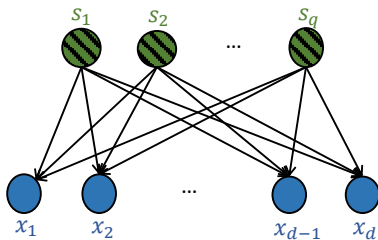
CPT: 8 different combination of values for 3 binary parents

Noisy-or: 4 parameters



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Noisy-OR Component Analyzer (NOCA)



Latent variables s : $(q + 1)$ -dimensions

$$s \in \{0,1\}^q, P(s_i|\pi_i) = \pi_i^{s_i}(1 - \pi_i)^{1-s_i}$$

$$\text{Loading Matrix: } \mathbf{p} = \{p_{ij}\}_{j=1,\dots,d}^{i=1,\dots,q}$$

$$q < d$$

Observed variables x : d -dimensions

$$x \in \{0,1\}^d$$

$$P(x) = \sum_{\{s\}} \left(\prod_{j=1}^d P(x_j|s) \right) \left(\prod_{i=1}^q P(s_i) \right)$$

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Why EM won't work?

- Take N iid samples (D-dimensional binary vectors)
- We will need:

- The joint distribution

$$P(\mathbf{x}, \mathbf{s}) = P(\mathbf{x}|\mathbf{s})P(\mathbf{s}) = P(\mathbf{s}) \prod_j \left(1 - \prod_{i=1}^K (1 - p_{ij})^{s_i} \right)^{x_j} \left(\prod_{i=1}^K (1 - p_{ij})^{s_i} \right)^{1-x_j}$$

Problem 1: not a product

- Joint over observables

$$P(\mathbf{x}) = \sum_{\mathbf{s}} P(\mathbf{x}, \mathbf{s}) = \sum_{\mathbf{s}} \left(\prod_j P(x_j|\mathbf{s}) \right) P(\mathbf{s})$$

Problem 2: summation over 2^K terms

Variational EM for NOCA

- Similar to what we did for CVQ, we simplify the distribution with a decomposable $Q(\mathbf{s})$

$$\begin{aligned} \log(P(\mathbf{x}|\theta)) &= \log\left(\prod_{n=1}^N P(x_n|\theta)\right) = \sum_{n=1}^N \log\left[\sum_{\{s\}} P(x_n, s_n|\theta)\right] \\ &= \sum_{n=1}^N \log\left[\sum_{\{s\}} P(x_n, s_n|\theta, q_n) \frac{Q(s_n)}{Q(s_n)}\right] \geq \sum_{n=1}^N \left[\sum_{\{s_n\}} E_{s_n} \log(P(x_n, s_n|\theta)) - E_{s_n} \log(Q(s_n))\right] \end{aligned}$$

- $\log(P(x_n, s_n|\theta, q_n))$ still can not be solved easily
- Noisy-Or is not in exponential family

Variational EM for NOCA

A further lower bound is required

- **Jensen's inequality:** $f(\mathbf{a} + \sum_j \mathbf{q}_j \mathbf{x}_j) \geq \sum_j \mathbf{q}_j f(\mathbf{a} + \mathbf{x}_j)$

$$P(x_j|s) = \left[1 - (1 - p_{0j}) \prod_{i=1}^q (1 - p_{ij})^{s_i} \right]^{x_j} \left[(1 - p_{0j}) \prod_{i=1}^q (1 - p_{ij})^{s_i} \right]^{(1-x_j)}$$

Set $\theta_{ij} = -\log(1 - p_{ij})$

$$P(x_j|s) = \exp \left[x_j \log \left(1 - \exp \left\{ -\theta_{0j} - \sum_{i=1}^q \theta_{ij} s_i \right\} \right) + (1 - x_j) \left(-\theta_{0j} - \sum_{i=1}^q \theta_{ij} s_i \right) \right]$$

$P(x_j|s)$ does not factorize for $x_j = 1$

$$\begin{aligned} P(x_j = 1|s) &= \exp[\log(1 - \exp\{-\theta_{0j} - \sum_{i=1}^q \theta_{ij} s_i\})] \\ &= \exp \left[\log \left(1 - \exp \left\{ -\theta_{0j} - \sum_{i=1}^q \theta_{ij} s_i \frac{q_j(i)}{q_j(i)} \right\} \right) \right] \geq \exp \left[\sum_{i=1}^q q_j(i) \log \left(1 - \exp \left\{ -\theta_{0j} - \frac{\theta_{ij} s_i}{q_j(i)} \right\} \right) \right] \\ &= \exp \left[\sum_{i=1}^q q_j(i) [s_i \log \left(1 - \exp \left\{ -\theta_{0j} - \frac{\theta_{ij}}{q_j(i)} \right\} \right) + (1 - s_i) \log(1 - \exp\{-\theta_{0j}\})] \right] \\ &= \prod_{i=1}^q \exp[q_j(i) s_i \log \left(1 - \exp \left\{ -\theta_{0j} - \frac{\theta_{ij}}{q_j(i)} \right\} \right) - \log(1 - \exp\{-\theta_{0j}\})] + q_j(i) \log(1 - \exp\{-\theta_{0j}\})] \end{aligned}$$

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Variational EM for NOCA

A further lower bound is required

$\log(P(x|\theta))$

$$\begin{aligned} &\geq \sum_{n=1}^N \left[\sum_{\{s_n\}} E_{s_n} \log P(x_n, s_n | \theta) - E_{s_n} \log Q(s_n) \right] \\ &\geq \sum_{n=1}^N \left[\sum_{\{s_n\}} E_{s_n} \log (\tilde{P}(x_n, s_n | \theta, q_n)) - E_{s_n} \log(Q(s_n)) \right] \\ &= \sum_{n=1}^N \left[\sum_{\{s_n\}} E_{s_n} \log (\tilde{P}(x_n | s_n, \theta, q_n) P(s_n | \theta)) - E_{s_n} \log(Q(s_n)) \right] \\ &= \sum_{n=1}^N \mathcal{F}_n(x_n, Q(s_n)) \\ &= \mathcal{F}(x, Q(s)) \end{aligned}$$

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Variational EM for NOCA

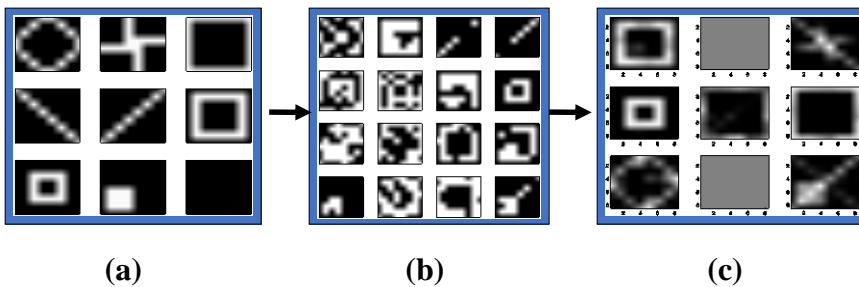
Parameters: $q_n, \theta_{ij}, \theta_{0j}$

• E-step: update q_n to optimize F_n

$$\bullet q_{nj}(i) \leftarrow \langle S_{ni} \rangle_{Q(S_n)} \frac{q_{nj}(i)}{\log(1 - e^{-\theta_{0j}})} \left[\log(1 - A^n(i, j)) - \frac{\theta_{ij}}{q_{nj}(i)} \frac{A^n(i, j)}{1 - A^n(i, j)} - \right]$$

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Structure Recovery Experiment



- a) Image patterns associated with hidden sources.
 b) Example images generated by the NOCA model
 c) Images recovered from source input.

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