

**CS 1675 Introduction to Machine Learning
Lecture 8**

Density estimation III

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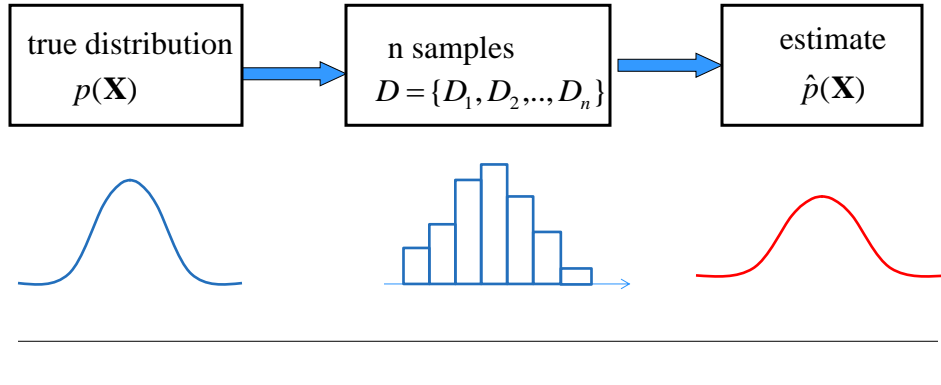
5329 Sennott Square

Parametric density estimation

Density estimation

Data: $D = \{D_1, D_2, \dots, D_n\}$
 $D_i = \mathbf{x}_i$ a vector of attribute values

Objective: estimate the model of the underlying probability distribution over variables \mathbf{X} , $p(\mathbf{X})$, using examples in D



ML Parameter estimation

Model $\hat{p}(\mathbf{X}) = p(\mathbf{X} | \Theta)$ **Data** $D = \{D_1, D_2, \dots, D_n\}$

• **Maximum likelihood (ML)** $\max_{\Theta} p(D | \Theta, \xi)$

– Find Θ that maximizes likelihood $p(D | \Theta, \xi)$

$$\begin{aligned}
 P(D | \Theta, \xi) &= P(D_1, D_2, \dots, D_n | \Theta, \xi) \\
 &= P(D_1 | \Theta, \xi) P(D_2 | \Theta, \xi) \dots P(D_n | \Theta, \xi) && \text{Independent examples} \\
 &= \prod_{i=1}^n P(D_i | \Theta, \xi)
 \end{aligned}$$

log-likelihood $\log p(D | \Theta, \xi) = \sum_{i=1}^n \log P(D_i | \Theta, \xi)$

$$\Theta_{ML} = \arg \max_{\Theta} p(D | \Theta, \xi) = \arg \max_{\Theta} \log p(D | \Theta, \xi)$$

Bayesian parameter estimation

Bayesian parameter estimation

- Uses the posterior distribution for parameters
- Posterior ‘covers’ all possible parameter values (and their “weights”)

$$p(\Theta | D, \xi) = \frac{p(D | \Theta, \xi) p(\Theta | \xi)}{p(D | \xi)}$$

Parameter posterior
Data Likelihood
Parameter prior

- How to use the posterior for modeling $p(\mathbf{X})$?

$$\hat{p}(\mathbf{X}) = p(\mathbf{X} | D) = \int_{\Theta} p(\mathbf{X} | \Theta) p(\Theta | D, \xi) d\Theta$$

Posterior Beta distribution



Prior Beta distribution

$$p(\theta | \xi) = \text{Beta}(\theta | \alpha_1, \alpha_2) = \frac{\Gamma(\alpha_1 + \alpha_2)}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \theta^{\alpha_1-1} (1-\theta)^{\alpha_2-1}$$

Why to use Beta distribution?

Beta distribution “fits” Bernoulli trials, it is called a **conjugate prior**

$$P(D | \theta, \xi) = \theta^{N_1} (1-\theta)^{N_2}$$

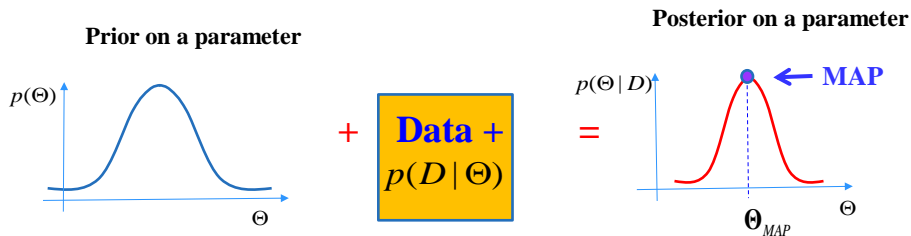
Posterior distribution is again a Beta distribution !!!!

$$\begin{aligned}
 p(\theta | D, \xi) &= \frac{P(D | \theta, \xi) \text{Beta}(\theta | \alpha_1, \alpha_2)}{P(D | \xi)} = \text{Beta}(\theta | \alpha_1 + N_1, \alpha_2 + N_2) \\
 &= \frac{\Gamma(\alpha_1 + \alpha_2 + N_1 + N_2)}{\Gamma(\alpha_1 + N_1)\Gamma(\alpha_2 + N_2)} \theta^{N_1 + \alpha_1 - 1} (1-\theta)^{N_2 + \alpha_2 - 1}
 \end{aligned}$$

Parameter estimation: MAP

- **Maximum a posteriori probability (MAP)**

$$\text{maximize } p(\Theta | D, \xi)$$



- **MAP**

- Yields: one set of parameters Θ_{MAP} (mode of the posterior)
- Approximation:

$$\hat{p}(\mathbf{X}) = p(\mathbf{X} | \Theta_{MAP})$$

Distribution models for random variables

Distribution models covered so far:

- **Bernoulli distribution**

- Model for binary random variables

$$P(x | \theta) = \theta^x (1 - \theta)^{(1-x)}$$

- **Binomial distribution**

- Model for order independent sets of binary outcomes

$$P(N_1 | N, \theta) = \binom{N}{N_1} \theta^{N_1} (1 - \theta)^{N - N_1}$$

- **Multinomial distribution**

- Model for order independent sets of k-nary outcomes

$$P(N_1, N_2, \dots, N_k | \theta, \xi) = \frac{N!}{N_1! N_2! \dots N_k!} \theta_1^{N_1} \theta_2^{N_2} \dots \theta_k^{N_k}$$

Distribution models for random variables

Models for other types of random variables:

- **Gaussian distribution**
 - Models of real-valued random variable
- **Gamma distribution:**
 - Models of random variables for positive real numbers
- **Exponential distribution**
 - Models of random variables for positive real numbers
- **Poisson distribution**
 - Models of random variables for nonnegative integers

Conjugate choices of priors for some these distributions:

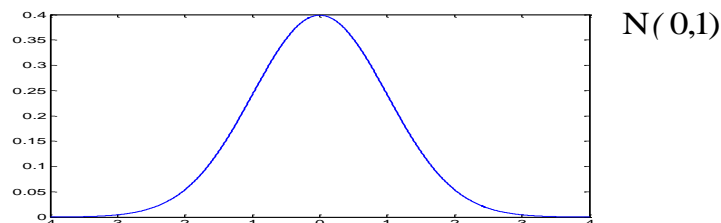
- **Exponential – Gamma**
- **Poisson – Inverse Gamma**
- **Gaussian - Gaussian (mean) and Wishart (covariance)**

Gaussian (normal) distribution

- **Gaussian:** $x \sim N(\mu, \sigma)$
- **Parameters:** μ - mean
 σ - standard deviation
- **Density function:**

$$p(x | \mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2\sigma^2}(x - \mu)^2\right]$$

- **Example:**



Parameter estimates

- **Loglikelihood**

$$l(D, \mu, \sigma) = \log \prod_{i=1}^n p(x_i | \mu, \sigma)$$

- **ML estimates of the mean and variance:**

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i \quad \hat{\sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2$$

- ML variance estimate is biased

$$E_n(\hat{\sigma}^2) = E_n\left(\frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2\right) = \frac{n-1}{n} \sigma^2 \neq \sigma^2$$

- **Unbiased estimate:**

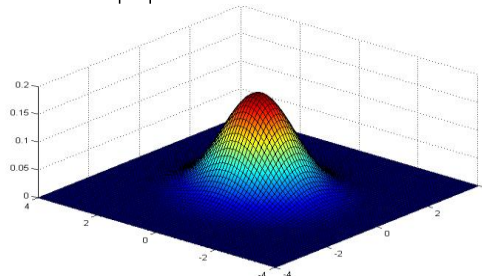
$$\hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu})^2$$

Multivariate normal distribution

- **Multivariate normal:** $\mathbf{x} \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$
- **Parameters:** $\boldsymbol{\mu}$ - mean
 $\boldsymbol{\Sigma}$ - covariance matrix
- **Density function:**

$$p(\mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{d/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right]$$

- **Example:**



Partitioned Gaussian Distributions

- **Multivariate Gaussian:**

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

- **Example:**

$$\mathbf{x} = \begin{pmatrix} \mathbf{x}_a \\ \mathbf{x}_b \end{pmatrix} \quad \boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{pmatrix} \quad \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_{aa} & \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba} & \boldsymbol{\Sigma}_{bb} \end{pmatrix}$$

$$\boldsymbol{\Lambda} \equiv \boldsymbol{\Sigma}^{-1} \quad \boldsymbol{\Lambda} = \begin{pmatrix} \boldsymbol{\Lambda}_{aa} & \boldsymbol{\Lambda}_{ab} \\ \boldsymbol{\Lambda}_{ba} & \boldsymbol{\Lambda}_{bb} \end{pmatrix}$$

Precision matrix

- **What are the distributions for marginals and conditionals?**

$$p(\mathbf{x}_a) \quad p(\mathbf{x}_a | \mathbf{x}_b)$$

Conditionals and Marginals

- **Conditional density:**

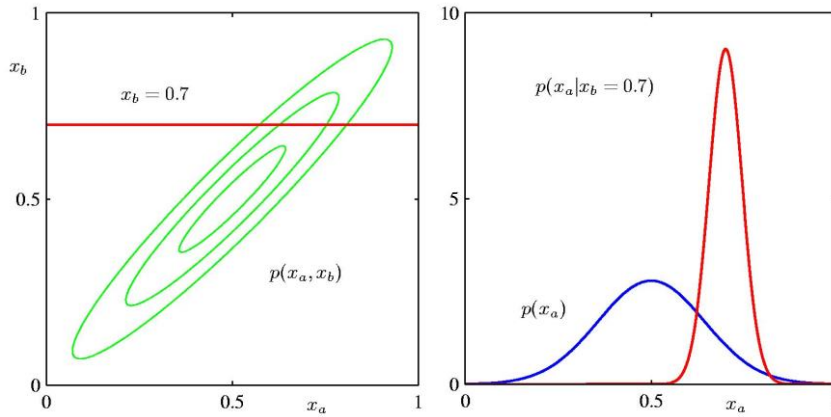
$$p(\mathbf{x}_a | \mathbf{x}_b) = \mathcal{N}(\mathbf{x}_a | \boldsymbol{\mu}_{a|b}, \boldsymbol{\Sigma}_{a|b})$$

$$\begin{aligned} \boldsymbol{\Sigma}_{a|b} &= \boldsymbol{\Lambda}_{aa}^{-1} = \boldsymbol{\Sigma}_{aa} - \boldsymbol{\Sigma}_{ab} \boldsymbol{\Sigma}_{bb}^{-1} \boldsymbol{\Sigma}_{ba} \\ \boldsymbol{\mu}_{a|b} &= \boldsymbol{\Sigma}_{a|b} \{ \boldsymbol{\Lambda}_{aa} \boldsymbol{\mu}_a - \boldsymbol{\Lambda}_{ab} (\mathbf{x}_b - \boldsymbol{\mu}_b) \} \\ &= \boldsymbol{\mu}_a - \boldsymbol{\Lambda}_{aa}^{-1} \boldsymbol{\Lambda}_{ab} (\mathbf{x}_b - \boldsymbol{\mu}_b) \\ &= \boldsymbol{\mu}_a + \boldsymbol{\Sigma}_{ab} \boldsymbol{\Sigma}_{bb}^{-1} (\mathbf{x}_b - \boldsymbol{\mu}_b) \end{aligned}$$

- **Marginal Density:**

$$\begin{aligned} p(\mathbf{x}_a) &= \int p(\mathbf{x}_a, \mathbf{x}_b) d\mathbf{x}_b \\ &= \mathcal{N}(\mathbf{x}_a | \boldsymbol{\mu}_a, \boldsymbol{\Sigma}_{aa}) \end{aligned}$$

Conditionals and Marginals



Parameter estimates

- **Loglikelihood**
$$l(D, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \log \prod_{i=1}^n p(\mathbf{x}_i | \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

- **ML estimates of the mean and covariances:**

$$\hat{\boldsymbol{\mu}} = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_i \quad \hat{\boldsymbol{\Sigma}} = \frac{1}{n} \sum_{i=1}^n (\mathbf{x}_i - \hat{\boldsymbol{\mu}})(\mathbf{x}_i - \hat{\boldsymbol{\mu}})^T$$

– Covariance estimate is biased

$$E_n(\hat{\boldsymbol{\Sigma}}) = E_n \left(\frac{1}{n} \sum_{i=1}^n (\mathbf{x}_i - \hat{\boldsymbol{\mu}})(\mathbf{x}_i - \hat{\boldsymbol{\mu}})^T \right) = \frac{n-1}{n} \boldsymbol{\Sigma} \neq \boldsymbol{\Sigma}$$

- **Unbiased estimate:**

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{n-1} \sum_{i=1}^n (\mathbf{x}_i - \hat{\boldsymbol{\mu}})(\mathbf{x}_i - \hat{\boldsymbol{\mu}})^T$$

Other distributions

Gamma distribution:

$$p(x | a, b) = \frac{1}{\Gamma(a)b^a} x^{a-1} e^{-\frac{x}{b}} \quad \text{for } x \in [0, \infty]$$

Exponential distribution:

- A special case of Gamma for $a=1$

$$p(x | b) = \left(\frac{1}{b}\right) e^{-\frac{x}{b}} \quad \text{for } x \in [0, \infty]$$

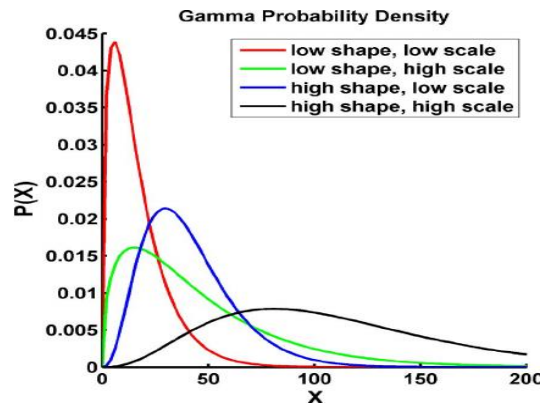
Poisson distribution:

$$p(x | \lambda) = \frac{e^{-\lambda} \lambda^x}{x!} \quad \text{for } x \in \{0, 1, 2, \dots\}$$

Gamma distribution

$$p(\lambda | a, b) = \frac{1}{\Gamma(a)b^a} \lambda^{a-1} e^{-\frac{\lambda}{b}} \quad \text{for } \lambda \in [0, \infty]$$

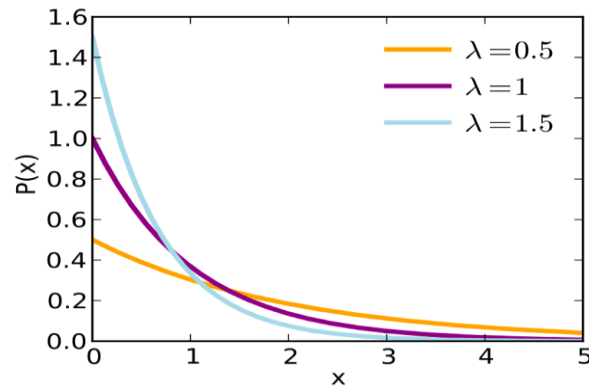
where a is the shape and b is a scale parameter



Exponential distribution

$$p(x | b) = \left(\frac{1}{b}\right) e^{-\frac{x}{b}} \quad \text{for } x \in [0, \infty]$$

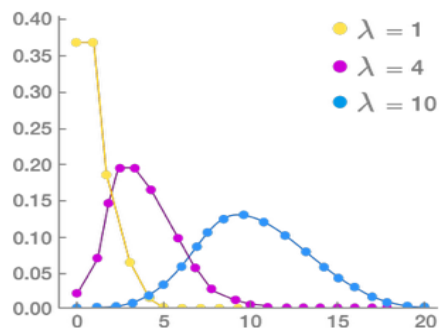
Alternative parameterization: $p(x | \lambda) = \lambda e^{-\lambda x}$
where $\lambda = 1/b$



Poisson distribution

Poisson distribution:

$$p(x | \lambda) = \frac{e^{-\lambda} \lambda^x}{x!} \quad \text{for } x \in \{0, 1, 2, \dots\}$$



Non-parametric density estimation

Nonparametric Density Estimation

- **Parametric distribution models** are:
 - restricted to specific functional forms, which may not always be suitable;
 - **Example:** modelling a multimodal distribution with a single, unimodal model.



- **Nonparametric approaches:**
 - Do not make any strong assumption about the overall shape of the distribution being modelled.
-

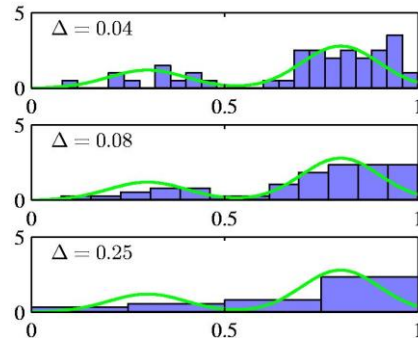
Nonparametric Methods

Histogram methods:

partition the data space into distinct bins with widths Δ_i and count the number of observations, n_i , in each bin.

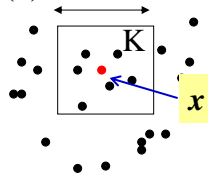
$$p_i = \frac{n_i}{N\Delta_i}$$

- Often, the same width is used for all bins, $\Delta_i = \Delta$.
- Δ acts as a smoothing parameter.
- Binning does not work well in the in a d -dimensional space,



Nonparametric Methods

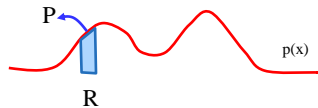
- Binning does not work well in a d -dimensional space,
 - M bins in each dimension will require M^d bins!
- **Solution:**
 - Build the estimates of $p(\mathbf{x})$ by considering the data points in D and how similar (or close) they are to \mathbf{x}
 - **Example: Parzen window**
 - As if we build a bin dynamically for \mathbf{x} for which we need to compute $p(\mathbf{x})$



Nonparametric Methods

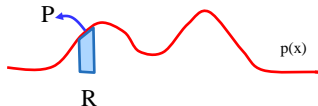
- Assume observations drawn from a density $p(x)$ and consider a small region R containing x such that

$$P = \int_R p(x) dx$$



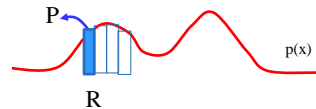
- The probability that K out of N observations lie inside R is $\text{Bin}(K, N, P)$ and if N is large

$$K \cong NP$$



If the volume of R , V , is sufficiently small, $p(x)$ is approximately constant over R and

$$P \cong p(x)V$$



Thus

$$p(x) = \frac{P}{V}$$

Putting things together we get:

$$p(x) = \frac{K}{NV}$$

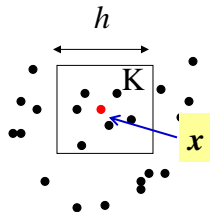
Nonparametric methods: kernel methods

Solution 1: Estimate the probability for x based on the fixed volume V built around x

$$p(x) = \frac{K}{NV}$$

- Fix V , estimate K from the data

Example: Parzen window



Nonparametric methods: kernel methods

Kernel Density Estimation:

- **Parzen window:** Let R be a hypercube centred on \mathbf{x} that defines the **kernel function:**

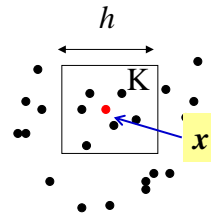
$$k\left(\frac{x-x_n}{h}\right) = \begin{cases} 1 & |x_i - x_{ni}| / h \leq 1/2 \\ 0 & \text{otherwise} \end{cases} \quad i = 1, \dots, D$$

- It follows that

$$K = \sum_{n=1}^N k\left(\frac{x-x_n}{h}\right)$$

- and hence

$$p(x) = \frac{K}{NV} = \frac{1}{Nh^D} \sum_{n=1}^N k\left(\frac{x-x_n}{h}\right)$$



Smooth kernels

To avoid discontinuities in $p(x)$ because of sharp boundaries we can use a **smooth kernel**, e.g. a Gaussian

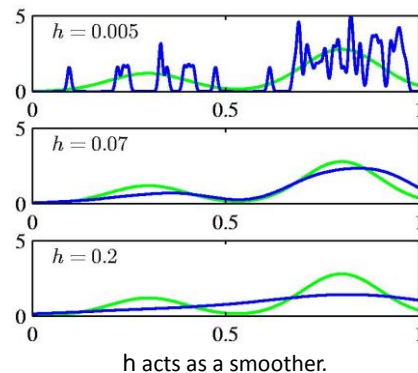
$$p(\mathbf{x}) = \frac{1}{N} \sum_{n=1}^N \frac{1}{(2\pi h^2)^{D/2}} \exp\left[-\frac{\|\mathbf{x} - \mathbf{x}_n\|^2}{2h^2}\right]$$

- Any kernel such that

$$k(\mathbf{u}) \geq 0$$

$$\int k(\mathbf{u}) d\mathbf{u} = 1$$

- will work.



Nonparametric Methods: kNN estimation

Solution 2: Estimate the probability for \mathbf{x} based on a fixed count \mathbf{K} for a variable volume \mathbf{V} built around \mathbf{x}

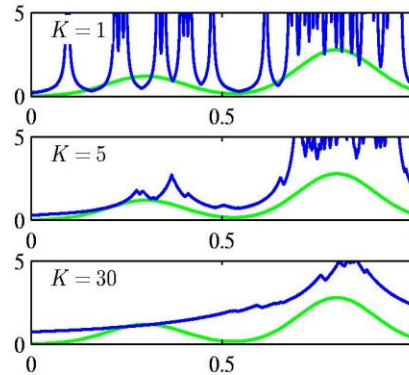
fix \mathbf{K} , estimate \mathbf{V} from the data

Nearest Neighbour Density Estimation:

Consider a hyper-sphere centred on \mathbf{x} and let it grow to a volume, V^* , that includes K of the given N data points.

Then

$$p(\mathbf{x}) \simeq \frac{K}{NV^*}.$$



K acts as a smoother

Nonparametric vs Parametric Methods

Nonparametric models:

- More flexibility – no density model is needed
- But require storing the entire dataset
- and the computation is performed with all data examples.

Parametric models:

- Once fitted, only parameters need to be stored
- They are much more efficient in terms of computation
- But the model needs to be picked in advance